

Bayesian Inference with Projected Densities

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Linear Inverse Problems

$$b = Ax + e$$

- observations $b \in \mathbb{R}^m$
- model/design/regressor matrix $A \in \mathbb{R}^{m \times n}$
- parameter vector $x \in \mathbb{R}^n$
- random measurement noise $e \in \mathbb{R}^m$

Bayesian approach

$$\underbrace{\pi(x|b)}_{\text{posterior}} \propto \underbrace{\pi(b|x)}_{\text{likelihood}} \underbrace{\pi(x)}_{\text{prior}}$$

What is a suitable noise model? What is a suitable prior?

Linear Inverse Problems (cont.)

Gaussian model

- Gaussian noise: $e \sim \mathcal{N}(0, \Sigma_e)$ with $\Sigma_e^{-1} = \lambda I$
- Gaussian prior: $x \sim \mathcal{N}(0, \Sigma_x)$ with $\Sigma_x^{-1} = \delta L^T L$
- Gaussian posterior

$$\pi(x|b) = \mathcal{N}(\hat{x}, \Sigma_{x|b}) \propto \exp\left(-\frac{\lambda}{2}\|Ax - b\|_2^2 - \frac{\delta}{2}\|Lx\|_2^2\right)$$

where $\Sigma_{x|b} = (\delta L^T L + \lambda A^T A)^{-1}$ and $\hat{x} = \lambda \Sigma_{x|b} A^T b$

- posterior mean is maximum a posteriori estimate
- Bayesian perspective: $\pi(x|b)$ is *the* solution

Sampling from the posterior distribution

Objective: compute samples from posterior distribution $\mathcal{N}(\hat{x}, \Sigma_{x|b})$

- Direct method: factorize $\Sigma_{x|b} = R^T R$ and compute

$$x = \hat{x} + R^T y, \quad y \sim \mathcal{N}(0, I)$$

- Randomize-then-optimize: generate $\hat{b} \sim \mathcal{N}(b, \lambda^{-1}I)$ and $\hat{c} \in \mathcal{N}(0, \delta^{-1}I)$ and compute

$$x = \operatorname{argmin}_u \left\{ \frac{\lambda}{2} \|Au - \hat{b}\|_2^2 + \frac{\delta}{2} \|Lu - \hat{c}\|_2^2 \right\}$$

Markov chain Monte Carlo and Plug & Play priors

- Unadjusted Langevin algorithm: samples from approximate posterior

$$x_{k+1} = x_k + \gamma \nabla \log \pi(y|x_k) + \gamma \nabla \log \pi(x_k) + \sqrt{2\gamma} u_{k+1}, \quad u_{k+1} \in \mathcal{N}(0, I)$$

- Apply Moreau–Yosida regularization if prior is nonsmooth:

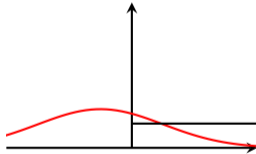
$$x_{k+1} = x_k + \gamma \nabla \log \pi(y|x_k) + \frac{\gamma}{\eta} [\text{prox}_U^\eta(x_k) - x_k] + \sqrt{2\gamma} u_{k+1}, \quad u_{k+1} \in \mathcal{N}(0, I)$$

- Replace $\eta^{-1}[\text{prox}_U^\eta(x_k) - x_k]$ by $\epsilon^{-1}[D_\epsilon(x_k) - x_k]$ where D_ϵ is a neural network (Laumont et al. 2022)

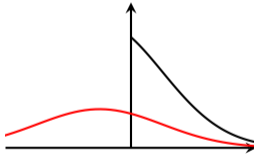
Additional prior information

Suppose x is known to be nonnegative, *i.e.*, $x \in \mathbb{R}_+^n$. What is a suitable prior?

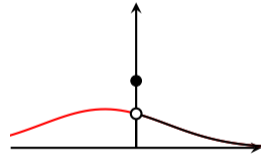
Uniform



Truncated Gaussian



Projected Gaussian



Sampling from projected Gaussian density

Oblique projection of $\pi_x = \mathcal{N}(\mu, \Sigma)$ onto a closed, convex set $C \subset \mathbb{R}^n$

$$z = \Pi_C^{\Sigma^{-1}}(x) = \operatorname{argmin}_{u \in C} \frac{1}{2} \|u - x\|_{\Sigma^{-1}}^2, \quad x \sim \mathcal{N}(\mu, \Sigma)$$

- z is well-defined since Σ is positive definite and $\Pi_C^{\Sigma^{-1}}$ is continuous

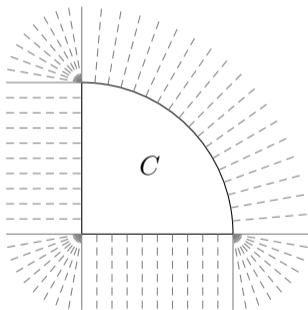
$$\mathbb{P}(\Pi_C^{\Sigma^{-1}}(x) \in E) = \mathbb{P}\left(x \in \left[\Pi_C^{\Sigma^{-1}}\right]^{-1}(E)\right)$$

- optimality condition for projection problem can be expressed as

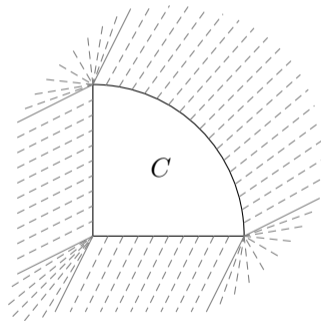
$$x \in z + \Sigma N_C(z)$$

Oblique projection

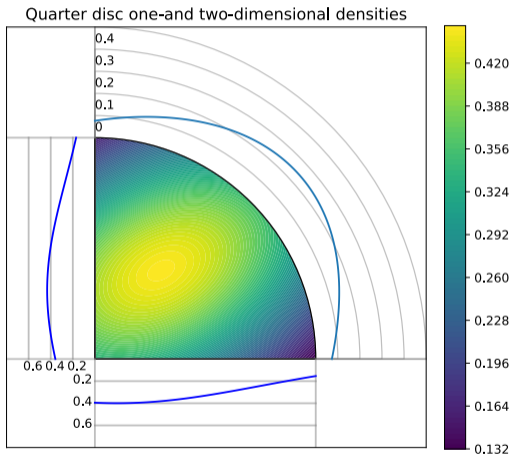
$$z + N_C(z)$$



$$z + \Sigma N_C(z)$$



Example



Non-empty, closed, convex set

- Positive probability on boundary and relative interior

$$\mathbb{P}(\Pi_C^{\Sigma^{-1}}(x) \in \mathbf{bd}(C)) > 0 \quad \text{and} \quad \mathbb{P}(\Pi_C^{\Sigma^{-1}}(x) \in \mathbf{relint}(C)) > 0$$

- Posterior mean is in the relative interior of C

$$\mathbb{E}[\Pi_C^{\Sigma^{-1}}(x)] \in \mathbf{relint}(C)$$

- MAP estimate is undefined

Polyhedral set

Suppose C is a polyhedral set and F is a face of C .

- For any measurable set $E \subseteq \mathbf{relint}(F)$, we have

$$\mathbb{P} \left(\Pi_C^{\Sigma^{-1}}(x) \in E \right) = \int_E \pi_F(z) dz$$

where $\pi_F(z)$ is the $\dim(F)$ -dimensional density

$$\pi_F(z) = \int_{\Sigma N_C(z)} \pi_x(z + v) dv.$$

- Projected density $\Pi_C^{\Sigma^{-1}}(x)$ is proportional to the density of x on the relative interior of any face of C .

Projected Gaussian posterior

Randomize-then-optimize: generate $\hat{b} \sim \mathcal{N}(b, \lambda^{-1}I)$ and $\hat{c} \in \mathcal{N}(0, \delta^{-1}I)$ and compute

$$x = \operatorname{argmin}_{u \in C} \left\{ \frac{\lambda}{2} \|Au - \hat{b}\|_2^2 + \frac{\delta}{2} \|Lu - \hat{c}\|_2^2 \right\}$$

- every sample requires the solution of randomized constrained LS problem
- posterior is well-defined, but what does the prior look like?
- suppose C is polyhedral and F is a face of C

$$\pi_{x,F}(u) \propto \pi_x(x_0 + Cu), \quad x_0 + Cu \in \mathbf{relint} F$$

Hierarchical Bayesian model

Hierarchical Bayesian model with hyper priors $\lambda \sim \Gamma(\alpha_\lambda, \beta_\lambda)$ and $\delta \sim \Gamma(\alpha_\delta, \beta_\delta)$

$$\pi(\lambda) \propto \lambda^{\alpha_\lambda - 1} \exp(-\beta_\lambda \lambda), \quad \text{for } \lambda > 0$$

$$\pi(\delta) \propto \delta^{\alpha_\delta - 1} \exp(-\beta_\delta \delta), \quad \text{for } \delta > 0$$

Full posterior

$$\begin{aligned} \pi_{x,\lambda,\delta|b}(x, \lambda, \delta) &\propto \lambda^{m/2 + \alpha_\lambda - 1} \delta^{n/2 + \alpha_\delta - 1} \\ &\times \exp\left(-\frac{\lambda}{2} \|Ax - b\|_2^2 - \frac{\delta}{2} \|Lx\|_2^2 - \beta_\lambda \lambda - \beta_\delta \delta\right), \end{aligned}$$

Polyhedral Cone Hierarchical Gibbs Sampler

Initialization: Choose $(x^0, \alpha_\lambda, \beta_\lambda, \alpha_\delta, \alpha_\lambda)$.

For $k = 1, 2, \dots$

Compute $(\lambda_k, \delta_k) \sim \pi_{\lambda, \delta | x, b}$ as follows:

$$\lambda_k \sim \Gamma \left(m/2 + \alpha_\lambda, \frac{1}{2} \|Ax_{k-1} - b\|_2^2 + \beta_\lambda \right)$$

$$\delta_k \sim \Gamma \left(\dim(F(x_{k-1}))/2 + \alpha_\delta, \frac{1}{2} \|Lx_{k-1}\|_2^2 + \beta_\delta \right)$$

Compute $x_k \sim \pi_{x | b, \lambda, \delta}$ by solving instance of:

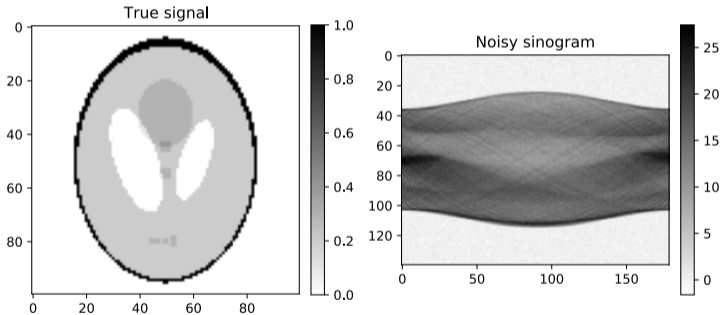
$$x_k = \operatorname{argmin}_{u \in C} \left\{ \frac{\lambda_k}{2} \|Au - \hat{b}\|_2^2 + \frac{\delta_k}{2} \|Lu - \hat{c}\|_2^2 \right\}$$

End.

Extension: convex regularization functions with polyhedral epigraph

Numerical example: CT reconstruction

$$b = Ax + e$$



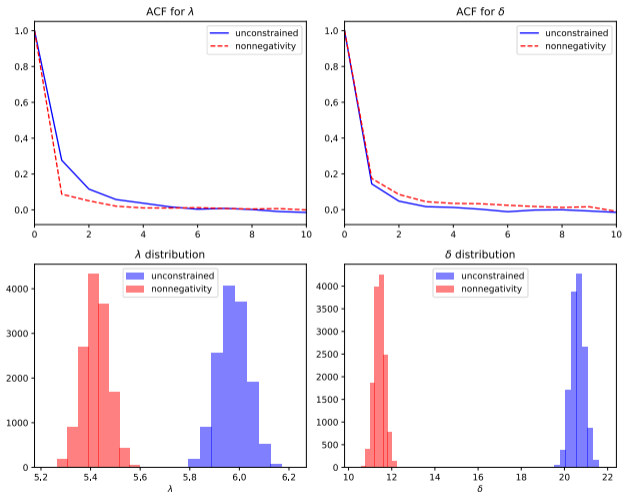
Numerical example: CT reconstruction (cont.)

Model

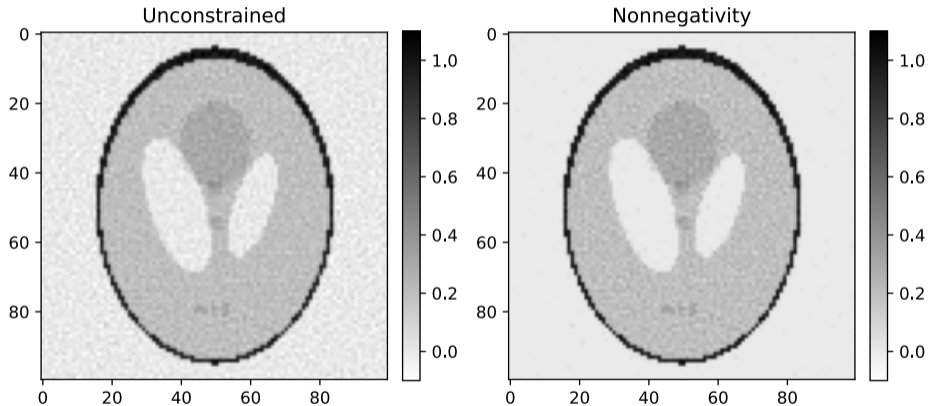
- L is first-order diff.
- $C = \mathbb{R}_+^n$

Gibbs sampler

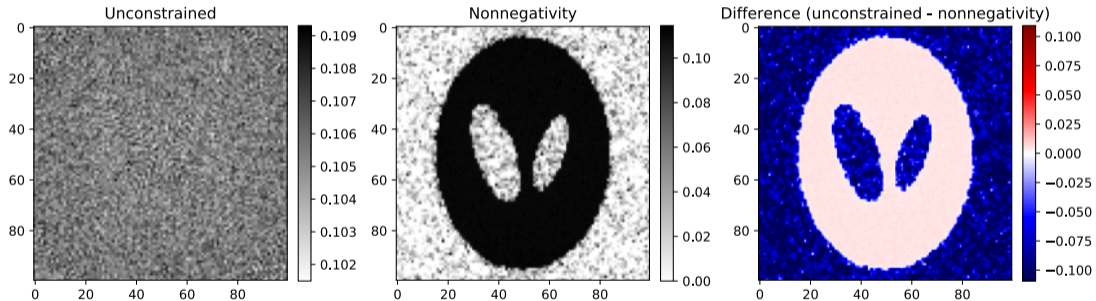
- 15000 samples (1000 burn-in)
- 1000 iterations with APG



Numerical example: CT reconstruction (cont.)











Numerical example: CT reconstruction (cont.)




Summary

- implicit prior defined through projected posterior
- positive probability on boundary
- Gibbs sampler when the set C is polyhedral
- preprint available on arXiv [2209.12481]

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